

Estimates of VaR for Itô Processes

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Risk Measures

- 1 VaR_α and Expected Shortfall.
- 2 Ruin Probability.
- 3 Expectation with respect an utility function.

VaR_α and Expected Shortfall

- If X is a r.v. $VaR_\alpha(X)$ is defined by

$$\begin{aligned} VaR_\alpha(X) &= \inf\{z \in R \mid P(X > z) < \alpha\} \\ &= \inf\{z \in R \mid P(X \leq z) \geq q\}, \end{aligned}$$

$$\alpha = 1 - q.$$

- Basel Accords regulate the credit institutions. One criteria is VaR_α $1 - \alpha = .99$, where X represent the loss. The institutions must be covered for losses inferior to this quantile.
- Expected Shortfall

$$E[X - VaR_\alpha \mid X > VaR_\alpha]$$

Difficulties

Estimation of very small probabilities. Different Methods:

- 1 Variance Reduction.
- 2 Bayesian
- 3 Extreme Value Theory. (Embrechts et al.)

For independent r.v. and for some kinds of dependence

Ruin Probability in Insurance

- Cramér-Lundberg Model:

$$X_t = x + ct - \sum_{i=1}^{N_t} Y_i$$

$$\text{Ruin Probability} = P[X_t < 0, \text{p. a. } t > 0] = \psi(x)$$

- Find x such the ruin probability $\psi(x) \leq q$.

$$\textit{Survival Probability} = \delta(x) = 1 - \psi(x)$$

$$\delta'(x) = \frac{\lambda}{c}\delta(x) - \frac{\lambda}{c} \int_0^x \delta(x-y)f(y)dy$$

If the r.v. Y_i admit Laplace transform

$$Ce^{-\theta x} \leq \psi(x) \leq e^{-\theta x}$$

Optimization w. r. Utility Function

- Utility function U , represents the risk aversion. (Concave Functions).

$$E[U(X)]$$

Stochastic Processes

If we have a stochastic processes X_t , $t \in [0, T]$ that represents the wealth or the price of a financial asset, what is the random variable we choose to calculate the Var?

- For each $t \in [0, T]$, X_t . Levy Processes. Important property independent and stationary increments.
- Some kind of Dynamic Var (Mc-Kean, Embrechts et al.) for time Series.

The General Model

Let B_t , $t \geq 0$ be a Brownian Motion and N_t , $t \geq 0$ a Poisson Process independent of B.M.

$$X_t = x + \int_0^t \sigma(s, \omega X_s) dB_s + \int_0^t b(s, \omega, X_s) ds + \sum_{i=1}^{N_t} \gamma_{T_i^-} Y_i, \quad t > 0,$$

- 1 Estimates of VaR and expected shortfall for processes without jumps.
- 2 Estimates of VaR and Ruin Probabilities for the General Model.

Joint work with Laurent Denis and Ana Meda.

$$dX_t = b(t, X_t)dt + \sigma(t, X_t)dW_t, \quad t > 0, \quad X_0 = m \quad (1)$$

We can take as our r.v X_t for $t > 0$ fixed.

- Estimates for the density (when it exists) of X_t can be used to obtain bounds for the VaR and the expected shortfall. In general, these are lower estimates. (Bally, V. Talay, D).

For $t \in R_+$ fixed

$$X_t^* = \sup_{0 \leq s \leq t} X_s$$

$$VaR_{m,\alpha}^t = \inf\{z \in R, P_m(X_t^* \geq z) \leq \alpha\}.$$

$$E[X_t^* - VaR_{m,\alpha}^t | X_t^* > VaR_{m,\alpha}^t] = E[X_t^* | X_t^* > VaR_{m,\alpha}^t] - VaR_{m,\alpha}^t$$

Main Idea

Obtain upper and lower bounds for the tail distribution of X_t^* , under some conditions of the coefficients and then

- 1 $VaR_{m,\alpha}^t$.
- 2 Expected shortfall.
- 3 The toy models.

Toy Models

1 Geometric Brownian Type

$$dX_t = b_t X_t dt + \sigma_t X_t dW_t.$$

2 Cox-Ingersoll Ross (Positive Process):

$$dX_t = (b - cX_t)dt + \sqrt{a^*} \sqrt{X_t} dW_t, \quad b, c, a \in \mathbb{Z}^+$$

3 Vasicek Type Model

$$dX_t = (\beta - \mu X_t)dt + \sigma dW_t, \quad \mu, \sigma \in \mathbb{R}^+, \beta \in \mathbb{R}.$$

Hypothesis (UB):

- 1 $b(t, z) \leq b^*, b^* \geq 0.$
- 2 $|\sigma(t, z)| \leq \sqrt{a^*}(z^+)^\gamma, a^* > 0, \gamma \in [0, 1).$

- 1 Geometric Brownian Type Model We will take

$$Y_t = \log(X_t), \quad dY_t = (b_t - \frac{\sigma_t^2}{2})dt + \sigma_t dW_t.$$

- 2 Cox-Ingersoll-Ross (Positive Process)

$$b(t, z) = b - cz^+ \leq b = b^*, \quad \sigma(t, z) = \sqrt{a^*} \sqrt{|z|}$$

- 3 Vasicek Type Model

$$Y_t = e^{\int_0^t \mu_s ds} X_t, \quad dY_t = \beta e^{\int_0^t \mu_s ds} dt + \sigma e^{\int_0^t \mu_s ds} dW_t$$

Lemma

Assume **(UB)** then for $t > 0$, $z \in \mathbb{R}$ we have

$$P_m(X_t^* \geq z) \leq 2\bar{\Phi} \left(\frac{(z - m - b^*t)^+}{\sqrt{a^*t}z^\gamma} \right).$$

where

$$\bar{\Phi}(x) = \frac{1}{\sqrt{2\pi}} \int_x^{+\infty} e^{-\frac{u^2}{2}} du.$$

Proof

If $z \geq m$, then we can assume

For all $y \geq z$, $a(y) = a(z)$.

$$P(X_t^* \geq z) \leq P\left(\sup_{0 \leq s \leq t} \int_0^s \sigma(X_u) dW_u \geq z - m - b^* t\right)$$

$$R_s = \int_0^s \sigma(X_u) dW_u, \quad R_s = \tilde{B}_{\langle R, R \rangle_s}, \quad \tilde{B}, B.M.$$

$$\begin{aligned} P(X_t^* \geq z) &\leq P\left(\sup_{0 \leq s \leq t} \tilde{B}_{\langle R, R \rangle_s} \geq z - m - b^* t\right) \\ &\leq P\left(\sup_{0 \leq u \leq a^* t z^{2\gamma}} \tilde{B}_u \geq z - m - b^* t\right) \\ &= 2\bar{\Phi}\left(\frac{(z - m - b^* t)^+}{\sqrt{a^* t} z^\gamma}\right). \end{aligned}$$

Upper Estimates

Theorem

Assume **(UB)**. For each $0 < t$,

$$\text{VaR}_{m,\alpha}^t \leq r,$$

where r is the unique root in $[b^*(t) + m, +\infty[$ of the equation

$$z - z^\gamma \sqrt{a^*(t)} \bar{\Phi}^{-1}(\alpha/2) - m - b^*(t) = 0.$$

Corollary

Assume (UB)

- ① *Geometric Brownian Type. Assume $b_s \leq b^*$, $a_* \leq \sigma_s \leq a^*$:*

$$\text{VaR}_{m,\alpha}^t \leq m e^{t(b^* - \frac{a_*}{2})^+ + \sqrt{a^*} t \bar{\Phi}(\frac{\alpha}{2})}$$

- ② *CIR*

$$\begin{aligned} \text{VaR}_{m,\alpha}^t &\leq m + b^* + \frac{1}{2} a^* (\bar{\Phi}^{-1})^2 (\alpha/2) \\ &+ \frac{1}{2} \bar{\Phi}^{-1} (\alpha/2) \sqrt{a^* (\bar{\Phi}^{-1})^2 (\alpha/2) + 4(m + b^*)} \end{aligned}$$

- ③ *Vasicek Type Model $\mu_t = \mu$*

$$\text{VaR}_{m,\alpha}^t \leq m + \beta e^{\mu t} + e^{\mu t} \sigma \sqrt{t} \bar{\Phi}^{-1} (\alpha/2)$$

Proposition

Assume (UB)

$$E(X_t^* | X_t^* \geq VaR_{m,\alpha}^t) \leq \frac{2}{\alpha} \int_{VaR_{m,\alpha}^t}^{\infty} \bar{\Phi} \left(\frac{(z - m - b^*)^+}{\sqrt{a^*} z^\gamma} \right) dz$$

Hypothesis LB

For all $z \in R$, $t \geq 0$,

- $b(t, z) \geq b_*$, $b_* \leq 0$.
- $\sigma(t, z) \geq a_*$, $a_* > 0$.

Lemma

Assume **(LB)**. Then for all $z \in \mathbb{R}$ y $t \geq 0$,

$$2\bar{\Phi}\left(\frac{(z - m - b_*t)^+}{\sqrt{a_*t}}\right) \leq P(X_t^* \geq z).$$

$$\bar{\Phi}(x) = \frac{1}{\sqrt{2\pi}} \int_x^{+\infty} e^{-\frac{u^2}{2}} du.$$

Theorem

Assume **(LB)**. Then, for all $t > 0$,

$$\text{VaR}_{m,\alpha}^t(X) \geq m + b_*t + \sqrt{a_*t}\bar{\Phi}^{-1}(\alpha/2).$$

Corollary

If $\gamma = 0$, i.e. σ bounded,

$$\text{VaR}_{m,\alpha}^{s,t}(X) \leq m + b^*(t - s) + \sqrt{a^*(t - s)}\bar{\Phi}^{-1}(\alpha/2).$$

X Geometric Brownian Type.

$$dX_t = b_t X_t dt + a_t X_t dW_t.$$

$$Y_t = \ln(X_t), \quad \text{for all } t \geq 0$$

$$Y_t = \ln(m) + \int_0^t \sigma_s dB_s + \int_0^t (b_s - \frac{\sigma_s^2}{2}) ds$$

If there exist constants $0 < a_* \leq a^*$ and $b_* \leq b^*$ such that $t \geq 0$

$$b_* \leq b_t \leq b^* \text{ and } a_* \leq \sigma_t^2 \leq a^*,$$

then for all $z \geq \ln m$ and $t \geq 0$

$$P(X_t^* \geq z) \leq 2\bar{\Phi} \left(\frac{(z - \ln m - t(b^* - \frac{a_*}{2})^+)^+}{\sqrt{a^* t}} \right).$$

$$2\bar{\Phi} \left(\frac{(z - \ln m - t(b_* - \frac{a^*}{2})^-)^+}{\sqrt{a_* t}} \right) \leq P(X_t^* \geq z).$$

Aplication

Consider a Russian option, on a Geometric Brownian type process. Let $t > 0$ be expiration time. The pay-off is given by

$$f_s = M_0 \bigvee \sup_{u \leq s} X_u.$$

If M_0 is fixed, $P_m(X_t^* \geq M_0)$ is the risk for the writter.

We can use the VaR to fix the value M_0 .

Dynamic VaR

$$X_{s,t}^* = \sup_{s \leq r \leq t} X_r$$

We want to measure the VaR of $X_{s,t}^*$ given that the process X exceeded VaR_α^s

This would help to know how risky the future is (up to time $t > s$), given that you already exceeded VaR in the past.

In a natural way we define

$$\widetilde{VaR}_\alpha^{s,t}(X) = \inf \left\{ z \in \mathbb{R}, P_X \left(X_{s,t}^* < z \mid X_s^* \geq VaR_{m,\alpha}^s \right) \geq q \right\}.$$

Theorem

Assume the process X has the Markov property, then

$$\widetilde{\text{VaR}}_{\alpha}^{s,t}(X) \leq r,$$

*where r is the unique root on $[b^*t + \text{VaR}_{m,\alpha}^s, +\infty[$ of the following equation:*

$$z - z^{\gamma} \sqrt{a^*t} \bar{\Phi}^{-1}(\alpha/2) - \text{VaR}_{m,\alpha}^s - b^*t = 0. \quad (2)$$

Proof

$$P_X(X_{s,t}^* \geq z | X_s^* \geq VaR_{X,\alpha}^s(X)) \leq \frac{P_X(X_{s,t}^* \geq z, X_s^* \geq VaR_{X,\alpha}^s(X))}{P_X(X_s^* \geq VaR_{X,\alpha}^s(X))}.$$

We now introduce

$$T = \inf \{u > 0, X_u = VaR_{X,\alpha}^s(X)\},$$

denote by μ the law of T under P_X . Then

$$\begin{aligned} P_X(X_{s,t}^* \geq z, X_s^* \geq VaR_{X,\alpha}^s(X)) &= \int_0^s P_X(X_{s,t}^* \geq z | T = r) \mu(dr) \\ &\leq \int_0^s P_X(X_{r,t}^* \geq z | T = r) \mu(dr) \\ &= \int_0^s P_{VaR_{X,\alpha}^s(X)}(X_{t-r}^* \geq z) \mu(dr) \end{aligned}$$

$$\begin{aligned}
& \int_0^s P_{VaR_{X,\alpha}^s(X)}(X_{t-r}^* \geq z) \mu(dr) \\
& \leq \int_0^s 2\bar{\Phi} \left(\frac{(z - VaR_{X,\alpha}^s(X) - b^*(t-r))^+}{\sqrt{a^*(t-r)} z^\gamma} \right) \mu(dr) \\
& \leq 2\bar{\Phi} \left(\frac{(z - VaR_{X,\alpha}^s(X) - b^*t)^+}{\sqrt{a^*t} z^\gamma} \right) P_X(T \leq s) \\
& = 2\bar{\Phi} \left(\frac{(z - VaR_{X,\alpha}^s(X) - b^*t)^+}{\sqrt{a^*t} z^\gamma} \right) P_X(X_s^* \geq VaR_{X,\alpha}^s(X)).
\end{aligned}$$

So,

$$P_X(X_{s,t}^* \geq z | X_s^* \geq VaR_{X,\alpha}^s(X)) \leq 2\bar{\Phi} \left(\frac{(z - VaR_{X,\alpha}^s(X) - b^*t)^+}{\sqrt{a^*t} z^\gamma} \right),$$

As previously, for $\gamma = 0$ or $\gamma = 1/2$ we are able to calculate this bound and this yields:

Corollary

(i) *If $\gamma = 0$, that is σ bounded, there is the following estimate:*

$$\widetilde{VaR}_{\alpha}^{s,t}(X) \leq VaR_{x,\alpha}^s(X) + b^*t + \sqrt{a^*t}\bar{\Phi}^{-1}(\alpha/2).$$

(ii) *If $\gamma = 1/2$ we have:*

$$\begin{aligned} \widetilde{VaR}_{\alpha}^{s,t}(X) &\leq VaR_{x,\alpha}^s(X) + b^*t + \frac{1}{2}a^*t(\phi^{-1})^2(\alpha/2) \\ &\quad + \frac{1}{2}\phi^{-1}(\alpha/2)\sqrt{a^*t(a^*t(\phi^{-1})^2(\alpha/2) + 4(VaR_{x,\alpha}^s(X) + b^*t))} \end{aligned}$$

Thanks

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